REVIEW



Protein intake and cancer: an umbrella review of systematic reviews for the evidence-based guideline of the German Nutrition Society

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Received: 19 September 2023 / Accepted: 23 March 2024 / Published online: 21 April 2024 © The Author(s) 2024

Abstract

Purpose It has been proposed that a higher habitual protein intake may increase cancer risk, possibly via upregulated insulinlike growth factor signalling. Since a systematic evaluation of human studies on protein intake and cancer risk based on a standardised assessment of systematic reviews (SRs) is lacking, we carried out an umbrella review of SRs on protein intake in relation to risks of different types of cancer.

Methods Following a pre-specified protocol (PROSPERO: CRD42018082395), we retrieved SRs on protein intake and cancer risk published before January 22th 2024, and assessed the methodological quality and outcome-specific certainty of the evidence using a modified version of AMSTAR 2 and NutriGrade, respectively. The overall certainty of evidence was rated according to predefined criteria.

Results Ten SRs were identified, of which eight included meta-analyses. Higher total protein intake was not associated with risks of breast, prostate, colorectal, ovarian, or pancreatic cancer incidence. The methodological quality of the included SRs ranged from *critically low* (kidney cancer), *low* (pancreatic, ovarian and prostate cancer) and *moderate* (breast and prostate cancer) to *high* (colorectal cancer). The outcome-specific certainty of the evidence underlying the reported findings on protein intake and cancer risk ranged from *very low* (pancreatic, ovarian and prostate cancer) to *low* (colorectal, ovarian, prostate, and breast cancer). Animal and plant protein intakes were not associated with cancer risks either at a *low* (breast and prostate cancer) or *very low* (pancreatic and prostate cancer) outcome-specific certainty of the evidence. Overall, the evidence for the lack of an association between protein intake and (i) colorectal cancer risk and (ii) breast cancer risk was rated as *possible*. By contrast, the evidence underlying the other reported results was rated as *insufficient*.

Conclusion The present findings suggest that higher total protein intake may not be associated with the risk of colorectal and breast cancer, while conclusions on protein intake in relation to risks of other types of cancer are restricted due to *insufficient* evidence.

Keywords Protein \cdot Umbrella review \cdot Evidence-based guideline \cdot Cancer

Abbreviations

| AMSTAR 2 | A Measurement Tool to Assess Systematic |
|----------|---|
| | Reviews 2 |
| FFQs | Food frequency questionnaires |
| IGF-1 | Insulin-like growth factor 1 |
| RCT | Randomised controlled trial |
| SR | Systematic review |
| WCRF | World Cancer Research Fund |
| | |

Introduction

The global incidence of cancer is on the increase. It is expected that cancer will become the most frequent cause of death in most countries worldwide during the twenty-first century [1, 2]. Despite advances in early detection and treatment, many cancer patients still have a lower life expectancy compared to persons without cancer. Moreover, the growing number of cancer survivors face major challenges with respect to tumour recurrence, comorbidities, psychosocial distress, and impairments to quality of life [3].

Extended author information available on the last page of the article

Around 40 percent of all cancers may be attributable to modifiable risk factors, and thus amenable to primary prevention [4–6]. Recent projections suggest that after smoking, which may explain 19.3% of all cancers [7], an unhealthy diet is the second most important cancer risk factor, with an attributable fraction of 7.8% due to low intakes of fruit, non-starchy vegetables, and fibre, and high intakes of salt, as well as red and processed meat [8]. The latter components of an unhealthy diet were derived from the World Cancer Research Fund's (WCRF) expert report on diet and cancer, a landmark compilation of systematic reviews (SRs) on associations between dietary factors and risks of different cancer types.

While increased risks of cancer and particularly colorectal cancer among regular consumers of red and processed meat have been explained by higher intakes of heterocyclic amines, polycyclic aromatic hydrocarbons, N-nitroso compounds, or haem iron [9], it has also been suggested that higher protein intake of animal origin—which may result from a high meat consumption—may be a dietary cancer risk factor [10]. Interestingly, laboratory-based mechanistic studies indicate that protein restriction may protect against cancer due to decreased systemic insulin-like growth factor 1 (IGF-1) signalling and a subsequent downregulation of intracellular mTor activation [11, 12].

While SRs on protein intake and risks of certain cancers have been published [13–22], a standardised assessment of the epidemiological evidence is lacking. Thus, we carried out an umbrella review of SRs of human studies on protein intake (overall, as well as animal and plant protein) and cancer risk (overall and by cancer type) based on standardised approaches to evaluate the quality of published SRs and to grade the certainty of the evidence. Our study is part of a series of umbrella reviews on protein intake and healthrelated outcomes, carried out as the basis of a new Protein Guideline by the German Nutrition Society [23].

Methods

We conducted an umbrella review (PROSPERO: CRD42018082395) following the methodological protocol published by Kroke et al. [23]. This protocol was developed as part of the evaluation of protein intake and various healthrelated outcomes and was also used for cancer. Two authors conducted all methodological steps independently. Any disagreements were resolved by discussion to achieve consensus.

Literature search

The systematic literature search was conducted in PubMed, Embase and Cochrane Database of Systematic Reviews for SRs published between 1st July/2009 and 22th January/2024. The date of 1st July 2009 originates from the decision to cover a ten-year period, i.e. the initial database search was conducted on 11th July 2019, and the last update was made on 22th January 2024 to capture the current status of research. Because this time period refers to the publication of the SRs, older primary literature is also taken into account. The search strategy is presented in the Supplementary Materials (S1). In addition to the systematic literature search in the three databases, SRs proposed by members of the guideline expert panel on protein of the German Nutrition Society were also included in the literature selection process.

Literature selection

Titles and/or abstracts of retrieved studies were screened according to pre-defined inclusion and exclusion criteria to identify potentially eligible studies [23]. The full-texts of potentially relevant publications were retrieved and assessed for eligibility. It was tolerated that some of the primary studies were incorporated more than once in different SRs. An overview of primary studies included in different SRs is shown in the Supplementary Materials (S2).

Publications were included if they met the following pre-specified criteria [23]: (i) SR evaluated the association between protein intake (total, animal-derived, plant-derived, from supplements) and cancer, (ii) population was the general adult population including older adults and recreational athletes, (iii) publication type was an SR with or without meta-analysis of prospective studies with human study participants, i. e. randomised controlled trials (RCTs), prospective cohort studies, case-cohort studies, or nested case-control studies. SRs also considering case-control studies were only included if prospective studies were predominant (> 50% of all studies), (iv) SR was written in English or German, and (v) SR was published between 07/2009 and 01/2024.

SRs of studies among children, adolescents, pregnant women, breastfeeding women, or top athletes were excluded, and so were individual studies, umbrella reviews, and SRs restricted to case–control studies, cross-sectional studies, or case studies. SRs of studies based on whole food approaches, in which protein intake was not specifically investigated, and studies on peptide and amino acid intakes were also excluded [23].

Data extraction

The following data from each included SR were extracted into a standardised Excel table: the first author's surname, year of publication, study type (e.g. SR with meta-analysis of RCTs), duration range of primary studies, study population, intervention/exposure(s), outcome(s), effect estimate(s) including 95% confidence intervals (CI), and heterogeneity estimate(s).

Methodological quality of SRs

To assess the methodological quality of the retrieved SRs, a modified version of the AMSTAR 2 (A Measurement Tool to Assess Systematic Reviews 2) tool [24] was used (Supplementary Materials (S3)). The modifications and the rationale for such modifications are described in detail in our methodological protocol [23]. This version of AMSTAR 2 contains 14 items evaluating the methodological quality of the SR. SRs were rated on a scale from high quality to critically low quality according to the existence of critical and non-critical methodological weaknesses. SRs rated as "critically low" by AMSTAR 2 were excluded from the rating of the overall certainty of evidence.

Outcome-specific certainty of evidence of SRs

The outcome-specific certainty of evidence of included SRs with and without meta-analysis was rated using the Nutri-Grade scoring tool [25] (Supplementary Materials (S4)). NutriGrade aims to assess the certainty of evidence of an association or effect between different dietary factors and outcomes, taking into account nutrition research-specific requirements not considered by other tools. This tool utilises a numerical scoring system and comprises seven items for SRs with meta-analysis of RCTs and eight items for SRs with meta-analysis of cohort studies. Based on the scoring system, four categories rate the potential outcome-specific certainty of evidence: high, moderate, low, and very low. The NutriGrade scoring tool was modified for the assessment of SRs without meta-analysis [23] (Supplementary Materials (S5)). We have adjusted the items related to meta-analyses: (i) precision: the confidence intervals were deleted, (ii) heterogeneity: this item was reduced to the question about consistency of the results, (iii) publication bias: this item was deleted, (iv) effect size: the RR/HR were deleted and (v) dose-response: this item was deleted. For SRs reporting more than one relevant outcome, each outcome was assessed separately. For SRs reporting more than one relevant exposure-outcome association, each exposureoutcome association was assessed separately.

Approaches to assess overall certainty of evidence and deriving conclusions

To reach a conclusion regarding protein intake and cancer, we proceeded in three steps. First, we assessed the methodological quality of retrieved SRs. Second, we used a scoring tool to assess the certainty of evidence of an association or effect between protein intake and the incidence of various cancer sites. Third, we rated the overall certainty of evidence separately for each relevant exposure-outcome association considering all relevant SRs and according to the framework outlined in the methodological protocol [23] and in Table 1. Briefly, the overall rating ranges from convincing, probable, possible to insufficient. At first, we assessed whether there is at least one SR with or without meta-analysis of prospective studies. If more than one SR with or without meta-analysis was available, all (convincing) or the majority (probable, possible) of the results must be consistent. Biological plausibility must be given in any case (direct or inverse association). In the final step, the results of the AMSTAR 2 and NutriGrade ratings were considered. Depending on the level of evidence, the SRs must have achieved a certain rating in both tools. If no SR is identified, or if the majority of SRs reached a very low outcome-specific certainty of evidence and/or low methodological quality, the overall certainty of evidence was considered insufficient. For this publication, two authors (KN and TK) made suggestions for rating the overall certainty of evidence. This rating was doublechecked by a staff member of the German Nutrition Society (NK) and thereafter reviewed by all co-authors. The final ratings of the overall certainty of evidence was approved by all authors.

Results

In total, 18,785 records were initially identified by literature search. An additional reference is based on the guideline expert panel's recommendation. This reference is an analysis of the WCRF that we included in the literature selection process. After the removal of duplicates, 15,066 records were screened based on title and abstract. We identified 163 potentially eligible records, and 10 were considered to be eligible with respect to inclusion and exclusion criteria [13–22]. The literature selection process is outlined in the flow diagram shown in Fig. 1. A list of excluded studies after full-text screening, including justifications for exclusion, is provided in Supplementary Materials (S6). Three SRs [27–29] were excluded from the evaluation because of a "critically low" AMSTAR 2 rating. The reason for exclusion was that Gathirua-Mwangi et al., Wu et al., and the WCRF report failed both to use a comprehensive literature search strategy and to provide an adequate risk of bias assessment. Additionally, the WCRF report failed to carry out an adequate investigation of publication bias.

Characteristics of the included SRs

Table 2 shows the characteristics of the 10 included SRs. None considered RCTs according to the search criteria defined by the respective authors, seven were SRs of only

| Overall certainty of evidence | Underlying criteria | Definition/explanation |
|-------------------------------------|---|--|
| Convincing | At least one SR with or without MA of prospective studies available If more than one SR with or without MA are available: all overall results must be consistent.¹ In case of a positive or negative association, biological plausibility is given All included SRs with or without MA must reach at least a "moderate" outcome-specific certainty of evidence²; in addition all included SRs must reach at least a methodological quality³ of "moderate" | There is high level of confidence that the true effect lies close to that of the estimate(s) of the effect |
| Probable | At least one SR with or without MA of prospective studies available If more than one SR with or without MA are available, the majority of overall results must be consistent.¹ In case of a positive or negative association, biological plausibility is given The majority⁴ of included SRs with or without MA must have reached at least a "moderate" certainty of evidence²; in addition all included SRs must reach at least a methodological quality³ of "moderate" | There is moderate confidence in the effect estimate(s): The true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different |
| Possible | • At least one SR with or without MA of prospective studies available | Confidence in the effect estimate(s) is limited: The true effect may be substantially different from the estimate |

Table 1 Grading the overall certainty of evidence according to methodological quality, outcome-specific certainty of evidence, biological p

certainty of evidence in a modified form according to the GRADE approach [23, 26]

OR

 1 Consistent = overall results of the SR have to be consistently either risk reducing or risk elevating or consistently showing no risk association

of the effect

estimate of effect

²Outcome-specific certainty of evidence refers to the NutriGrade rating

• If more than one SR with or without MA are available, the

• The majority⁴ of included SRs with or without MA must reach at least a "low" certainty of evidence²; in addition the majority⁴ of all included SRs must reach at least a methodo-

• The majority⁴ of included SRs with or without MA reach a

"very low" certainty of evidence²; in addition the majority of all included SRs reach a methodological quality³ of "low"

majority of overall results must be consistent.1 · In case of a positive or negative association, biological

plausibility is given

· No SR is available

logical quality³ of "moderate"

³Methodological quality refers the AMSTAR 2 rating; SRs graded as "critically low" by AMSTAR 2 are not considered

⁴Majority: > 50% of the included SRs

cohort studies [13-15, 17, 18, 20, 21], two were SRs of cohort and case control studies [16, 19], and one was an SR of cohort and nested case-control studies [22]. Two out of the ten retrieved SRs did not include a meta-analysis [17, 22]. Dose-response analyses were conducted in one of the eight meta-analyses [14].

The included SRs comprised 32 primary studies and investigated the following outcomes: prostate cancer [14, 17, 19], ovarian cancer [18, 20], colorectal cancer [21], breast cancer [13, 22], pancreatic cancer [16] and overall cancer incidence [15]. Eight SRs included assessments of total protein in relation to cancer risks [14, 16–22], and separate assessments of plant protein and animal protein were reported in four SRs [14, 16, 19, 22]. Two SRs addressed soy protein and cancer risk [13, 15]. Two SRs investigated dairy protein in relation to cancer risk [14, 17]. The study duration of the included primary studies ranged from 3.5 to 18 years for cohort studies. The SRs on colorectal, pancreatic and overall cancer incidence included both sexes, while the SRs on prostate cancer included only male participants and the SRs on breast and

There is very little confidence in the effect estimate (s): The true effect is likely to be substantially different from the

Insufficient

MA meta-analysis, SR systematic review



Fig. 1 Flow diagram on systematic reviews included

ovarian cancer only included female participants. The age of the participants ranged from 18 to 92 years, and participants were free of cancer at baseline.

Type and range of protein intake

The type and amount of protein intake was not provided in five SRs [14–17, 22]. For the other five SRs the reported total protein intake ranged between 70 and 129 g/d vs. 0–84 g/d, the range of animal protein intake was 41–80 g/d vs. 9–47 g/d, and the range of plant protein intake was 28–47 g/d vs. 13–29 g/d for cohort studies, although exact minimum and maximum values for protein intakes were

not available from all original studies included. Of note, food frequency questionnaires (FFQ) were used to estimate protein intakes in the vast majority of the original studies included in the SRs.

Methodological quality

Overall scores of AMSTAR 2 for each included SR are reported in Table 2. Supplementary Materials (S7) provide a more detailed overview showing the assessments of each individual item. Methodological quality was rated high for one SR [21], moderate for five SRs [13–15, 17, 22] and low for four SRs [16, 18–20].

| A) Systematic rev | A) Systematic reviews with meta-analysis | S | | | | | | | |
|------------------------|--|---|--------------------------------|--|-----------------|--|-----------------------------|--|--------------------|
| Author, year | Study type, study period | Study population | Exposition | Protein intake | Outcome | Effect estimates RR (95% CI) | Heterogeneity estimators | Heterogeneity NutriGrade rating estimators | AMSTAR 2 rating |
| Shin 2023 [13] | - Published before 12/2022 - Follow-up: 7.4 yr | - Only females - age NP | High vs. low protein intake | > 12.8 w. ≤4.87 g/d Dietary assessment methods NP | Breast cancer | | | | Moderate |
| | 1 cohort study | n = 73,223 participants n = 592 cases | Soy protein | | | 0.89 (0.67,1.17) ¹ | $I^{2} = 0\%$ | NA | |
| Alzahrani 2022 [14] | - Published before 10/2021 - Follow-up: 3.5–18 yr | - Only males - Aged 26–92 yr | High vs. low protein intake | Protein intake NP Dietary assessment methods: selfreported FFQ, selfreported food record, recall interview | Prostate cancer | | | | Moderate |
| | 8 cohort studies | n = 214,276 participants pants n = 12,567 cases | Total protein | 90.0–128.7 g/d vs. 38.0–82.0 g/d | | $0.99 (0.92, 1.07)^2$ | $I^2 = 12.8\%$ P = 0.33 | Low: 4.9 | |
| | 7 cohort studies | n = 204,060 partici- pants n = 7314 cases | | | | Dose-response analysis: 0.99 (0.98, 1.01) per 20 g/d increase P _{non-lineutiy} =0.51 | $P^2 = 0\%$ P = 0.43 | | |
| | 6 cohort studies | n = 324,197 participants n = 26,207 cases | Animal protein | 62.8–80 g/d <i>vs.</i> 8.5–47 g/d | | 0.99 (0.95, 1.04) | $I^2 = 0\%$ P = 0.94 | Low: 5.9 | |
| | 5 cohort studies | n = 313,981 partici- pants n = cases | | | | Dose-response analysis: 1.00 (0.98, 1.01) per 20 g/d increase P _{non-lineutiy} =0.34 | $I^2 = 25.5\%$ P = 0.25 | | |

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| Table 2 (continued) | nued) | | | | | | | | |
|---------------------|--|---|--------------------------------|---|-----------------------------|--|-----------------------------|-------------------|--------------------|
| A) Systematic re | A) Systematic reviews with meta-analysis | is | | | | | | | |
| Author, year | Study type, study period | Study population | Exposition | Protein intake | Outcome | Effect estimates RR (95% CI) | Heterogeneity estimators | NutriGrade rating | AMSTAR 2 rating |
| | 6 cohort studies | n = 329,082 partici- pants n = 26,137 cases | Plant protein | 27. <u>9-47.0 g/d vs.</u> 13.0-29.0 g/d | | 1.01 (0.96, 1.06) | $I^2 = 0\%$ P = 0.63 | Low: 4.9 | |
| | 4 cohort studies | n = 304,866 partici- pants n = 20,704 cases | | | | Dose-response analysis: 1.01 (0.98, 1.04) per 20 g/d increase P _{non-lineutiy} =0.81 | $I^2 = 0\%$ P = 0.57 | | |
| | 4 cohort studies | n = 167,489 partici- pants n = 9,864 cases | Dairy protein | 20.5–27.0 g/d vs. 6.3–10.0 g/d | | 1.08 (1.00, 1.16) | $I^2 = 38.1\%$ P = 0.16 | Moderate: 6.0 | |
| | 3 cohort studies | n = 157,273 partici- pants n = 4611 cases | | | | Dose-responseanalysis: 1.10 (1.02, 1.20) per 20 g/d increase "Dairy protein intakes from 30 g/d to higher amounts were associ- ated with a higher risk of prostate cancer, while this association for the intakes of < 30 g/d was not significant." Pnon-linearity = 0.02 | P=0.17 | | |
| Fan 2022 [15] | - Published before 09/2021 - Follow-up: 5.2–14 yr | - Both sexes - Aged 40–74 yr | High vs. low protein intake | Protein intake NP Dietary assessment methods: 77-item validated FFQ, 81-item validated FFQ, 45-item validated FFQ, 165-item validated FFQ | Overall cancer incidence | | | | Moderate |
| | 6 cohort studies | n = 349,158 partici- pants n = 1,970 cases | Soy protein | | | 0.95 (0.71, 1.28) ¹ | $I^2 = 68.9\%$ P = 0.007 | Low: 4.3 | |

| A) Systematic reviews with meta-analysis | reviews with meta-analysi | 0 | | | | | | | |
|--|--|---|--------------------------------|---|-------------------|-----------------------------------|-----------------------------|--|--------------------|
| Author, year | Study type, study period | Study population | Exposition | Protein intake | Outcome | Effect estimates RR (95% CI) | Heterogeneity estimators | Heterogeneity NutriGrade rating estimators | AMSTAR 2 rating |
| Zhang 2022 [16] | - Published before 10/2019 - Follow-up: NP | - Both sexes - Aged 34–89 yr | High vs. low protein intake | Protein intake NP Dietary assessment methods: FFQ, validated FFQ | Pancreatic cancer | | | | Low |
| | 2 cohort studies | n = 77,156 participants pants n = 217 cases | Total protein | | | 0.98 (0.63, 1.54) ¹ | $I^2 = 0\%$ P = 0.758 | Very low: 1.5 | |
| | 2 cohort studies and 1 case-control study | n = 78,134 partici- pants n = 543 cases | Animal protein | | | $1.37 (0.93, 2.01)^{1}$ | $I^2 = 32.7\%$ P = 0.226 | Very low: 2.5 | |
| | 2 cohort studies and 1 case–control study | n = 78,134 partici- pants n = 543 cases | Plant protein | | | $0.78 (0.54, 1.14)^{1}$ | $I^2 = 11.5\%$ P = 0.323 | Very low: 2.5 | |
| Khodavandi 2020 [18] | - Published before 11/2019 - Follow-up: 8.1–16 yr | - Only females - Aged 25–84 yr | High vs. low protein intake | > 83.5 g/d vs. <72.6 g/d g/d Dietary assessment methods: haseline question- naire with 55-food naire with 55-food items, 126-item FPQ, center-spe- cific dietary questionnaires (EPIC), FFQ (block) | Ovarian cancer | | | | Low |
| | 4 cohort studies | <i>n</i> = 464,643 participants <i>n</i> = 1,585 cases | Total protein | | | 0.96 (0.88, 1.06) ² | $I^2 = 0\%$ P = 0.741 | Low: 5.0 | |
| Mao 2018 [19] | - Published before 09/2017 - Follow-up: 3.5-18 yr | - Only males - Aged 40–92 yr | High vs. low protein intake | Total protein: 90–121 g/d vs. 0–82 g/d Animal protein: 41–80 g/d vs. ≤47 g/d 35–47 g/d vs. ≤29 g/d Dictary assessment methods NP | Prostate cancer | | | | Low |
| | 8 cohort studies | n = 265,067 partici- pants n = 5,860 cases | Total protein | | | $1.080 (0.964, 1.209)^{1}$ | $I^2 = 0\%$ P = 0.670 | Very low: 2.9 | |
| | 4 cohort studies and 1 case-control study | n = 220,916 partici- pants n = 8,826 cases | Animal protein | | | 1.001 (0.917, 1.092) ¹ | $I^2 = 0\%$ P = 0.891 | Very low: 3.5 | |

Table 2 (continued)

| Author, vear | A) Systematic reviews with meta-analysis | S | | | | | | | | | |
|----------------------------|--|---|--------------------------------|--------------------------------|--|---|---------------------|---|---|--|----------------------|
| | Study type, study period | Study population | Exposition | Protein intake | intake | Outcome | Effect estir CI) | Effect estimates RR (95% CI) | Heterogeneity estimators | Heterogeneity NutriGrade rating estimators | g AMSTAR 2 rating |
| | 4 cohort studies and 1 case–control study | n = 234,462 partici- pants n = 8,937 cases | Plant protein | | | | 0.986 (0.9 | 0.986 (0.904, 1.076) ¹ | $P^2 = 0\%$ P = 0.556 | Very low: 3.5 | |
| Pang 2018 [20] | - Published until 06/2018 - Follow-up: NP | - Only females - Aged 18–87 yr | High vs. low protein intake | | > 89.9 g/d vs. <72.6 g/d Dietary assessment methods NP | > 89.9 g/d vs. < 72.6 Ovarian cancer g/d Dietary assessment methods NP | | | | | Low |
| | 2 cohort studies | n = 63,275 participants pants n = 210 cases | Total protein | _ | | | 0.903 (0.6 | 0.903 (0.679, 1.201) ¹ | $I^2 = 27.9\%$ P = 0.239 | Very low: 2.5 | |
| Lai 2017 [21] | - Published until 12/2016 - Follow-up: NP | - Both sexes - Aged 20–89 yr | High w. low protein intake | | 115.9 g/d vs. 83.8 g/day dietary assessment methods: Semiquantitative FFQ, FFQ, | Colorectal cancer | 2 | | | | High |
| | 3 cohort studies | n = 680 cases number of partici- pants: NP | Total protein | | | | 0.939 (0.7 | 0.939 (0.730, 1.209) ¹ | $I^2 = 0\%$ P = 0.980 | Low: 4.0 | |
| B) Systematic | B) Systematic reviews without meta-analysis | analysis | | | | | | | | | |
| Author, year | Study type, study period | tudy Study population | | Exposition | Prote | Protein intake | Outcome | Results | Nut | NutriGrade rating A | AMSTAR 2 rating |
| Ubago-Guisado 2021 [17] | do - Published until 05/2021 - Follow-up: 8.7 yr | ntil - Only males - Age NP 8.7 yr | | High vs. low protein intake | | Protein intake NP Dietary assessment methods NP | Prostate cancer | er | | Z | Moderate |
| | 1 cohort study | | rtici- | Total protein | | | | "Total protein intake was not positively associated with increased risk (HR 1.17 (0.96, 1.44)" | "Total protein intake NA was not positively associated with increased risk (HR 1.17 (0.96, 1.44)" | | |
| | 1 cohort study | | rtici- | Dairy protein | | | | "Protein from dairy foods was signifi- cantly associated with an increased risk (HR 1.22 (1.07, 1.41)" "An increment of 35 g/d dairy protein | n dairy NA ilgnifi- ciated reased 22)" tu of 35 rotein | | |
| | | | | | | | | was associated with an HR of 1.32 (1.01, 1.72)" | ated 0f 1.32 | | |

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| B) Systematic reviews without meta-analysis | s without meta-analysis | | | | | | | |
|---|---|--|--------------------------------|---|---------------|---|-----------------------------------|-------------|
| Author, year | Study type, study period | Study population | Exposition | Protein intake | Outcome | Results | NutriGrade rating AMSTAR 2 rating | AR 2 rating |
| Pedersen 2013 [22] | - Published between 01/2000 and 12/2011 - Follow-up: 5.5–18 yr | - Only females - Healthy - Aged 30–70 yr | High vs. low protein intake | Protein intake np Dietary assessment methods: 61-item fiq over the past year, 107-item fiq over the previ- ous year | Breast cancer | | Moderate | ą |
| | 1 cohort study and 1 nested case-con- trol study | <i>n</i> = 88,647 participants <i>n</i> = 214 controls <i>n</i> = 4,163 cases | Total protein | | | None of the two studies found statistically signifi- cant associations | Low: 4.0 | |
| | 1 cohort study and 1 nested case-con- trol study | <i>n</i> = 88,647 participants <i>n</i> = 214 controls <i>n</i> = 4,163 cases | Animal protein | | | Between total, animal or plant protein and breast cancer risk | Low: 4.0 | |
| | 1 cohort study and 1 nested case-con- trol study | - n = 88,647 participants - n = 214 controls - n = 4,163 cases | Plant protein | | | | Low: 4.0 | |

|y|5 /al, d day(s 5 2 Disto AMSTAR 2 A Measurement 1 year(s), HR hazard ratio ¹random effect model

²fixed effect model

Outcome-specific certainty of the evidence

None of the included SRs found a statistically significant (p < 0.05) association between protein intake and cancer risk and risk estimates comparing highest versus lowest intake categories were largely close to 1. One exception is a statistically significant positive association that was observed between dairy protein intake and prostate cancer risk [14, 17].

Overall scores of NutriGrade for each SR are summarised in Table 2. Briefly, out of the 17 NutriGrade ratings, seven were very low, nine were low, and one was moderate. Supplementary Materials (S8) provides a more detailed account showing the assessments of each individual Nutri-Grade item.

Rating of the overall certainty of the evidence

According to our pre-specified criteria [23], the overall evidence for the lack of an association between total protein intake and colorectal cancer and between total, animal and plant protein intake and breast cancer was rated as *possible*. By contrast, the evidence regarding animal and plant protein consumption in relation to colorectal cancer risk was rated as *insufficient*, in view of a *low* methodological quality of the underlying studies. For the same reason, the overall evidence on associations between protein intake (total, animalderived, and plant-derived) and ovarian and prostate cancer was rated as *insufficient*.

Discussion

In the present umbrella review, 10 SRs, of which eight also included meta-analyses, were identified to derive the available evidence on the association between dietary protein intake and cancer risks, considering colorectal cancer, breast cancer, ovarian cancer, prostate cancer, pancreatic cancer and overall cancer incidence. Overall, no association between dietary protein intake and risk of cancer was observed. However, the evidence for this lack of an association was only rated as *possible* with regard to total protein intake and colorectal and breast cancer risk, while the evidence for associations between protein intake and the evidence for other cancer types was rated as *insufficient*.

Previous umbrella reviews have been conducted in the field of diet and cancer risk [30–34], but not all of them explicitly included protein intake. An umbrella review on the role of diet on colorectal cancer incidence included dietary protein intake as exposure and referred to the SR by Lai et al. 2017 (which is also included in our umbrella review) considering the association between dietary protein intake

as association with non-significant evidence with respect to colorectal cancer incidence [30].

The studies included in the here reviewed SRs and metaanalyses assessed dietary intake mainly by FFQs, which are prone to misclassification bias and limited in determining absolute intake values, but have been shown to be appropriate to rank study participants according to their dietary intakes [35, 36]. The use of different nutrient databases for the calculation of protein intake from food intake may have introduced additional heterogeneity and misclassification. However, we included here SRs that are largely based on cohort studies, in which the misclassification due to the dietary assessment is often non-differential since the exposure is measured before the outcome [37]. Nevertheless, such non-differential misclassification is particularly expected for protein intake estimated from FFQs as compared to estimation of fat and carbohydrate intake [38] and may have led to severe attenuation of estimated protein intake-cancer relationships impairing the detection of small associations [39]. In addition, analyses on animal versus plant protein intake may be especially prone to residual confounding due to associations with other lifestyle factors that may be related to cancer risk. Finally, it cannot be excluded that other cooccurring ingredients in protein-rich foods may play a role in the observed associations.

It has been suggested that high protein intake may be related to cancer risk through alterations in the IGF-1 axis. IGF-1 physiologically regulates cellular proliferation, differentiation, and apoptosis [40], but may also exert mitogenic and anti-apoptotic effects that promote carcinogenesis [41, 42]. High circulating IGF-1 concentrations have been associated with cancer risk [43], recently supported by Mendelian randomisation studies, particularly for breast, colorectal and prostate cancer [44–47], while no significant associations were observed for ovarian cancer [45], and no corresponding evidence is available for pancreatic cancer. There is evidence that high protein intake is associated with higher circulating IGF-1 from intervention studies [48], as well as from observational studies [49–57], although some smaller observational studies did not observe significant associations [58–62]. There is also evidence that higher animal [49], and particularly dairy protein intake [49, 55, 56] is associated with higher circulating IGF-1 [60]. However, it is questionable whether the impact of protein intake on IGF-1 concentrations is sufficient to observe altered cancer risks on a population level. For example, in a recent investigation within the UK Biobank comparing the highest (mean 89.8 g/day) versus lowest (mean 68.1 g/day) quintile of total protein intake was associated with 1.72 nmol/L higher IGF-1 concentrations [57]. However, based on the Mendelian Randomisation analysis on cancer in the UK Biobank [45], such a difference would translate to only 1% differences in risk of colorectal (OR per 5.7 nmol/L higher IGF-1 was 1.11, 95%

CI 1.01–1.22) or prostate cancer (OR per 5.7 nmol/L higher IGF-1 was RR 1.10; 95% CI 1.01–1.21).

Besides prospective studies investigating protein intake on the macronutrient level in relation to cancer risk, numerous studies have been conducted on the association between protein-rich foods and cancer risk. For example, dairy and especially milk intake has been related to moderately higher risk of prostate cancer, and it has been hypothesised that the underlying mechanism is through dairy protein intake increasing circulating IGF-1 [63]. However, a recent Mendelian randomisation study found only limited evidence for an association between genetically determined higher milk intake (proxied through a genetic variant associated with lactase persistence in adulthood) and risk of prostate cancer [64]. In the two SRs with meta-analysis on prostate cancer included in the present umbrella review [14, 19], no associations were observed for total protein, plant or animal protein intake. In the more recent SR with meta-analysis on prostate cancer, a significant positive association between dairy protein intake and prostate cancer was observed. In contrast to the observed positive association with prostate cancer, dairy intake has been consistently associated with lower risk of colorectal cancer, recently confirmed by a Mendelian randomisation study [64], but the hypothesised mechanism is here not through dairy protein, but through dairy calcium [65, 66]. Another protein source that has been consistently associated with risk of colorectal cancer is meat intake [67]. However, since positive associations were observed with red and processed meat intake and no or inverse associations with white meat intake [68], the underlying mechanism may be unrelated to protein. In line with this suggestion, red and processed meat intake have been shown to be only weakly associated with IGF-1 concentration [69]. Heme iron and compounds formed during the preparation such as heterocyclic amines, polycyclic aromatic hydrocarbons, as well as nitrate and N-nitrosocompounds, have been suggested here as underlying mechanisms for the observed positive associations between red and processed meat and prostate cancer risk [70, 71]. The SR with meta-analysis [21] included in this umbrella review did not observe an association between total protein intake and risk of colorectal cancer summarising data from three cohort studies, and no association for animal (5 case-control or cohort studies) or plant (4 case-control or cohort studies) protein intake. Dairy intake has also been related to lower risk of breast cancer [72, 73], but a recent pooled analysis of cohort studies comprising more than one million women did not find associations [74], and also the evidence from Mendelian randomisation studies was inconclusive [64].

Protein sources from plant origin such as legumes [75] have been inconclusively related to cancer risk, while soy

intake has been related to lower risk of breast cancer [76], although the suggested underlying mechanism is ascribed to phytoestrogens, and thus probably unrelated to protein per se. The SRs with meta-analysis on soy protein intake included here did not observe an association with overall cancer incidence [15] or breast cancer risk [13].

We acknowledge that through alterations in the IGF-1 system there is a plausible biological mechanism that may mediate positive associations between protein intake and cancer risk although the quantitative importance may be limited. At the same time, other nutrients that co-occur with protein in food sources, such as calcium in milk, which may reduce colorectal cancer risk, and heme iron and heterocyclic amines in red meat and N-nitroso-compounds in processed meat, which may increase colorectal cancer risk, as well as phytoestrogens in soy, which may reduce breast cancer risk, may mask potential associations with overall protein intake. This may explain why on the macronutrient level protein intake has not been consistently related to cancer risk.

Strengths of the current umbrella review include the comprehensive literature search, as well as the assessment of the methodological quality of the included SRs with AMSTAR 2 and the rating of their outcome-specific certainty of evidence with NutriGrade. We applied Nutri-Grade instead of the GRADE approach (Grading of Recommendations, Assessment, Development and Evaluation) because an important novelty of NutriGrade (published in 2016) was the modified classification for meta-analyses of RCTs and cohort studies compared with the traditional GRADE approach (initially classifying RCTs with an initial high score and cohort studies with a low score) [77]. We are aware that in the meantime the GRADE approach was amended (adjustments published in 2019, but after the guideline methodology was established in 2017) in a way that cohort studies can now also be assigned an initially high score, when risk of bias tools such as ROBINS-I are used [78]. A general limitation of this umbrella review is that all included SRs were based on observational studies where dietary assessment was usually performed through FFQs, thereby limiting the ability to draw conclusions regarding absolute intakes and to detect small associations due to measurement error. Only a limited number of SRs were available for inclusion in the present umbrella review, and for certain types of cancer (e.g. lung cancer, kidney cancer, bladder cancer), no SRs were available. The scarcity of SRs on the topic may also be related to the fact that protein intake is more rarely investigated in prospective cohort studies than food intake. A further limitation is the limited generalizability of results given that the prospective cohort studies included in the here reviewed SRs are usually not representative of the general population. In addition, we cannot exclude that the associations observed in the included SRs were subject to residual confounding due to insufficient confounder adjustment in the original studies.

Conclusion

The present umbrella review concludes that there is no association between total protein intake and risk of colorectal cancer and between total, animal and plant protein intake and risk of breast cancer, at *possible* certainty of the evidence. The certainty of the evidence is *insufficient* with regard to intakes of animal- and plant-derived protein in relation to colorectal cancer. Moreover, the certainty of evidence for the lack of associations between protein intake (total, animal-derived and plant-derived) and risks of breast, ovarian, and prostate cancer, as well as overall cancer incidence from the identified SRs, is also *insufficient*. Thus, unless stronger evidence from prospective studies becomes available, our overall finding that protein intake may not be associated with cancer risk needs to be interpreted with caution.

Supplementary Information The online version contains supplementary material available at https://doi.org/10.1007/s00394-024-03380-4.

Acknowledgements We would like to thank all members of the guideline expert panel on protein of the German Nutrition Society for their contributions to the methodological approach and specifically to the manuscript on cancer. Not all of the remarks and suggestions during expert panel meetings and at previous versions of the manuscript met the criteria for co-authorship, but were very helpful in shaping the final version manuscript. The following scientists deserve special thanks: Jürgen M. Bauer, Heike A. Bischoff-Ferrari, Heiner Boeing, Thomas Remer and Dorothee Volkert.

Author contributions AL, AMA, AS, JH and NK conducted the systematic literature search, literature selection, data extraction and assessment of methodological quality and outcome-specific certainty of evidence. KN and TK evaluated the evidence and graded the overall certainty of evidence, which was finalised after discussion with all guideline panel members. KN, TK, AS and NK prepared the manuscript. AB, SEg, SE, AK, SL, SLo, MBS, LS, RS, GIS, BW and AZ provided critical feedback for revision. All authors read and approved the manuscript.

Funding This research was funded by the German Federal Ministry of Food and Agriculture. The funder had no role in the decisions regarding data collection, analyses, interpretation of data, in the writing of the report or in the decision to submit the article for publication.

Data availability Not applicable.

Declarations

Conflict of interest A list of possible conflicts of interest is provided in Supplementary Material S9.

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